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Binary hidden Markov models and varieties

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Abstract. This paper closely examines HMMs in which all the hidden random variables are binary. Its main contributions are (1) a birational parametrization for every such HMM, with an explicit inverse for recovering the hidden parameters in terms of observables, (2) a semialgebraic model membership test for every such HMM, and (3) minimal defining equations for the 4-node fully binary model, comprising 21 quadrics and 29 cubics, which were computed using Gröbner bases in the cumulant coordinates of Sturmfels and Zwiernik. The new model parameters in (1) are rationally identifiable in the sense of Sullivant, Garcia-Puente, and Spielvogel, and each model's Zariski closure is therefore a rational projective variety of dimension 5. Gröbner basis computations for the model and its graph are found to be considerably faster using these parameters. In the case of two hidden states, item (2) supersedes a previous algorithm of Schönhuth which is only generically defined, and the defining equations (3) yield new invariants for HMMs of all lengths ≥ 4 . Such invariants have been used successfully in model selection problems in phylogenetics, and one can hope for similar applications in the case of HMMs.

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1. Introduction

The present work is motivated primarily by the problems of model selection and parameter identifiability, viewed from the perspective of algebraic geometry. By beginning with the simplest hidden Markov models (HMMs) — those where all hidden nodes are binary — the hope is that eventually a very precise geometric understanding of HMMs can be attained that provides insight into these central problems. Indeed, most questions about this case are answered by reducing to the case where the visible nodes are also binary. The history of this and related problems has two main branches of historical lineage: that of hidden Markov models, and that of algebraic statistics.

Hidden Markov models were developed as statistical models in a series of papers by Leonard E. Baum and others beginning with Baum and Petrie [4], after the description

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by Stratonovich [20] of the "forward-backward" algorithm that would be used for HMM parameter estimation. HMMs have been used extensively in natural language processing and speech recognition since the development of DRAGON by Baker [1]. As well, since Krogh et al. [13] used HMM for gene finding in the DNA of in E. coli bacteria, they have had many applications in genomics and biological sequence alignment; see also [24]. Now, HMM parameter estimation is built into the measurement of so many kinds of time-series data that it would be gratuitous to enumerate them. However, the methods of algebraic statistics are not so old, and the algebraic geometry of these models is far from fully explored. They are hence an important early example for the theory to investigate.

Algebraic statistics is the application of commutative algebra and algebraic geometry to the study of statistical models, especially those models involving non-linear relations between parameters and observables. It was first described at length in the monograph Algebraic Statistics by Pistone, Riccomagno, and Wynn [16]. Subsequent introductions to the subject include Algebraic Statistics for Computation Biology by Pachter and Sturmfels [15], and Lectures in Algebraic Statistics by Drton, Sturmfels, and Sullivant [9]. Also notable is Algebraic Geometry and Statistical Learning Theory by Watanabe [23], for its focus on the problem of model selection from data.

To the problem of model selection, the algebraic analogue is implicitization, i.e., finding polynomial defining equations for the Zariski closures of binary hidden Markov models. Such polynomials are called invariants of the model: if a polynomial f is equal to a constant c at every point on the model (i.e. does not vary with the model parameters), then we encode this equation by calling f-c "an invariant". Model selection and implicitization are more than simply analogous; polynomial invariants have been used successfully in model selection by Casanellas and Fernandez-Sanchez [6] and Eriksson [10] for phylogenetic trees.

Invariants have been difficult to classify for hidden Markov models, perhaps due to the high codimension of the models. Bray and Morton [5] found many invariants using linear algebra, but did not exhibit any generating sets of invariants, and in fact their search was actually for invariants of a model that was slightly modified from the HMM proper. Schönhuth [17] found a large family of invariants arising as minors of certain non-abelian Hankel matrices, and was able to verify that such invariants generate the ideal of the 3-node binary HMM, the simplest non-degenerate HMM. However, this seemed not to be the case for models with $n \geq 4$ nodes: Schönhuth reported on a computation of J. Hauenstein which verified numerically that the 4-node model was not cut out by the Hankel minors.

In Section 3, we will make use of moment and cumulant coordinates as exposited in [21], as well as a new coordinate system on the parameter space, to find explicit defining equations for the 4-node binary HMM. The shortest quadric and cubic equations are fairly simple; to give the reader a visual sense, they look like this:

$$\begin{split} g_{2,1} &= m_{23} m_{13} - m_2 m_{134} - m_{13} m_{12} + m_1 m_{124} \\ g_{3,1} &= m_{12}^3 - 2 m_1 m_{12} m_{123} + m_{\emptyset} m_{123}^2 + m_1^2 m_{1234} - m_{\emptyset} m_{12} m_{1234} \end{split}$$

Here each m is a moment of the observed probability distribution. These equations are not generated by Schönhuth's Hankel minors, and so provide a finer test for membership to any binary HMM of length $n \geq 4$ after marginalizing to any 4 consecutive nodes.

To the problem of parameter identifiability, the algebraic analogue is the generic or global injectivity or finiteness of a map of varieties that parametrizes the model, or in the case of identifying a single parameter, constancy of the parameter on the fibers of the parameterization. Sullivant et al. [22] provide an excellent discussion of this topic in the context of identifying causal effects; see also [14] for a striking application to identification for ODE models in the biosciences.

In Section 4, for the purpose of parameter identification in binary hidden Markov models, we express the parametrization of a binary HMM as the composition of a dominant and generically finite monomial map \mathfrak{q} and a birationally invertible map ψ . An explicit inverse to ψ is given, which allows for the easy recovery of hidden parameters in terms of observables. The components of the monomial map are *identifiable combinations* in the sense of Meshkat et al. [14]. The formulae for recovering the hidden parameters are fairly simple when exhibited in a particular order, corresponding to a particular triangular set of generators in a union of lexicographic Gröbner bases for the model ideal. To show their simplicity, the most complicated recovery formula looks like this:

$$u = \frac{m_1 m_3 - m_2^2 + m_{23} - m_{12}}{2(m_3 - m_2)}$$

As a corollary, in Section 4.3 we find that the fibers of ϕ_n are generically zero-dimensional, consisting of two points which are equivalent under a "hidden label swapping" operation.

Section 5 describes describes how the parametrization of every fully binary HMM, or "BHMM", can be factored through a particular 9-dimensional variety called a trace variety, which is the invariant theory quotient of the space of triples of 2×2 matrices under a simultaneous conjugation action by SL_2 . As a quotient, the trace variety is not defined inside any particular ambient space. However, its coordinate ring, a trace algebra, was found by Sibirskii [18] to be generated by 10 elements, which means we can embed the trace variety, and hence all BHMMs simultaneously, in \mathbb{C}^{10} . We prove the main results of Section 4 in the coordinates of this embedding. As a byproduct of this approach, in section Section 5.6 we find that the Zariski closures of all BHMMs with $n \geq 3$ are birational to each other.

Finally, Section 6 explores some applications of our results, including model membership testing, classification of identifiable parameters, a new grading on HMMs that can be used to find low-degree invariants, the geometry of equilibrium BHMMs, and HMMs with more than two visible states.

2. Definitions

Important note: In this paper, we will work mostly with BHMMs — HMMs in which both the hidden and visible nodes are all binary — because, as will be explained in Section 2.3, all our results will generalize to allow ≥ 2 visible states by reducing to this case.

Throughout, we will be referring to binary hidden Markov processes, distributions, maps, models, varieties, and ideals. Each of these terms is used with a distinct meaning,

and effort is made to keep their usages consistent and separate.

2.1. Binary Hidden Markov processes and distributions

A binary hidden Markov process is a statistical process which generates random binary sequences. It is based on the simpler notion of a binary (and not hidden) Markov *chain* process.

Definition 1. A Binary Hidden Markov process will comprise 5 data: π , T, E, and (H_t, V_t) . The pair (H_t, V_t) denotes a jointly random sequence $(H_1, V_1, H_2, V_2, ...)$ of binary variables, also respectively called hidden nodes and visible nodes, with range $\{0, 1\}$. Often a bound n on the (discrete) time index t is also given. The joint distribution of the nodes is specified by the following:

- A row vector $\pi = (\pi_0, \pi_1)$, called the initial distribution, which specifies a probability distribution on the first hidden node H_1 by $\Pr(H_1 = i) = \pi_i$;
- A matrix $T = \begin{bmatrix} T_{00} & T_{01} \\ T_{10} & T_{11} \end{bmatrix}$, called the transition matrix, which specifies conditional "transition" probabilities by the formula $\Pr(H_t = j \mid H_{t-1} = i) = T_{ij}$, read as the probability of "transitioning from hidden state i to hidden state j".*
- A matrix $E = \begin{bmatrix} E_{00} & E_{01} \\ E_{10} & E_{11} \end{bmatrix}$, called the emission matrix, which specifies conditional "emission" probabilities by the formula $\Pr(V_t = j \mid H_t = i) = E_{ij}$, read as the probability that "hidden state i emits the visible state j".

To be precise, the parameter vector $\theta = (\pi, T, E)$ determines a probability distribution on the set of sequences of pairs $((H_1, V_1) \dots (H_n, V_n)) \in (\{0, 1\}^2)^n$, or if no bound n is specified, a compatible sequence of such distributions as n grows. In applications, only the joint distribution on the visible nodes $(V_1, \dots, V_n) \in \{0, 1\}^n$ is observed, and is called the *observed distribution*. This distribution is given by marginalizing (summing) over the possible hidden states of a BHM process:

$$\Pr(V = v \mid \theta = (\pi, T, E)) = \sum_{h \in \{0,1\}^n} \Pr(h, v \mid \pi, T, E) = \sum_{h \in \{0,1\}^n} \Pr(h \mid \pi, T) \Pr(v \mid h, E)$$
$$= \sum_{h \in \{0,1\}^n} \pi_{h_1} E_{h_1, v_1} \prod_{i=2}^n T_{h_{i-1}h_i} E_{h_i, v_i}$$
(1)

Definition 2. A Binary Hidden Markov distribution is a probability distribution on sequences $v \in \{0,1\}^n$ of jointly random binary variables (V_1, \ldots, V_n) which arises as the observed distribution of some BHM process according to (1).

^{*[17]} uses T for different matrices, which I will later denote by P.

As we will see in Section 4.1, different processes (π, T, E, H_t, V_t) can give rise to the same observed distribution on the V_t , for example by permuting the labels of the hidden variables, or by other relations among the parameters.

Those already familiar with Markov models in some form may note that:

- The data (π, T, H_t) alone specify what is ordinarily called a binary Markov *chain* process on the nodes H_t . In the applications we have in mind, these nodes are unobserved variables.
- The matrices T and E are assumed to be *stationary*, meaning that they are not allowed to vary with the "time index" t of (H_t, V_t) .
- The distribution π is not assumed to be at equilibrium, i.e. we do not assume that $\pi T = \pi$. This allows for more diverse applications.

Remark 1. The term "stationary" is sometimes also used for a process that is at equilibrium; we will reserve the term "stationary" for the constancy of matrices T,E over time.

2.2. Binary Hidden Markov maps, models, varieties, and ideals

Statistical processes come in families defined by allowing their parameters to vary, and in short, the set of probability distributions that can arise from the processes in a given family is called a *statistical model*. The Zariski closure of such a model in an appropriate complex space is an algebraic variety, and the geometry of this variety carries information about the purely algebraic properties of the model.

In a binary hidden Markov process, π , T, and E must be *stochastic matrices*, i.e. each of their rows must consist of non-negative reals which sum to 1, since these rows are probability distributions. We denote by $\Theta_{\rm st}$ the set of such triples (π, T, E) , which is isometric to the 5-dimensional cube $(\Delta_1)^5$. We call $\Theta_{\rm st}$ the space of *stochastic parameters*. It is helpful to also consider the larger space of triples (π, T, E) where the matrices can have arbitrary complex entries with row sums of 1. We write $\Theta_{\mathbb{C}}$ for this larger space, which is equal to complex Zariski closure of $\Theta_{\rm st}$, and call is the space of *complex parameters*.

We will not simply replace Θ_{st} by $\Theta_{\mathbb{C}}$ for convenience, as has sometimes been done in algebraic phylogenetics. For the ring of polynomial functions on these spaces, we write

$$\mathbb{C}[\theta] := \mathbb{C}[\pi_j, T_{ij}, E_{ij}] / \left(1 = \sum_j \pi_j = \sum_j T_{ij} = \sum_j E_{ij} \text{ for } i = 0, 1\right)$$

so as to make the identification $\Theta_{st} \subseteq \Theta_{\mathbb{C}} = \operatorname{Spec} \mathbb{C}[\theta]$. Here Spec denotes the spectrum of a ring; see [7] for this and other fundamentals of algebraic geometry.

Now we a fix a length |v| = n for our binary sequences v, and write

$$R_{p,n} := \mathbb{C}[p_v \mid v \in \{0,1\}^n] \qquad \qquad \mathbb{C}_p^{2^n} := \operatorname{Spec}(R_{p,n})$$

$$\overline{R}_{p,n} := R_{p,n} / (1 - \sum_{|v|=n} p_v) \qquad \qquad \mathbb{C}_p^{2^n-1} := \operatorname{Spec}(\overline{R}_{p,n})$$

$$\mathbb{P}_p^{2^n-1} := \operatorname{Proj}(R_{p,n})$$

We will often have occasion to consider the natural inclusions,

$$\iota_n: \mathbb{C}_p^{2^n-1} \hookrightarrow \mathbb{C}_p^{2^n}$$
 $\bar{\iota}_n: \mathbb{C}_p^{2^n-1} \hookrightarrow \mathbb{P}_p^{2^n-1}$

Convention 2.1. Complex spaces such as \mathbb{C}^{2^n} will usually be decorated with a subscript to indicate the intended coordinates to be used on that space, like the p in $\mathbb{C}_p^{2^n}$ above. Likewise, a ring will usually be denoted by R with some subscripts to indicate its generators.

Definition 3. For $n \geq 3$,

• The Binary Hidden Markov map or modeling map on n nodes is the map $\phi_{\text{BHMM}(n)}$, or simply ϕ_n , given by given by (1), i.e.

$$\phi_n:\Theta_{\mathbb{C}}\to\mathbb{C}_p^{2^n-1},$$

$$\phi_n^{\#}(p_v) := \sum_{h \in \{0,1\}^n} \pi_{h_1} E_{h_1,v_1} \prod_{i=2}^n T_{h_{i-1}h_i} E_{h_i,v_i}$$

The word "model" is also frequently used for the map ϕ_n . This is a very reasonable usage of the term, but I reserve "model" for the image of the allowed parameter values:

• BHMM(n), the **Binary Hidden Markov** model on n nodes, is the image

$$\bar{\iota}_n \phi_n \left(\Theta_{\mathrm{st}}\right) \subseteq \mathbb{P}_p^{2^n - 1},$$

of the stochastic parameter space Θ_{st} , i.e., the set of observed distributions which can arise from some BHM process, considered as a subset of $\mathbb{P}_p^{2^n-1}$ via $\bar{\iota}_n$. Being the continuous image of the classically compact cube $\Theta_{st} \simeq \Delta_1^5$, BHMM(n) is also classically compact and hence classically closed.

- $\overline{\operatorname{BHMM}}(n)$, the Binary Hidden Markov variety on n nodes, is the Zariski closure of $\operatorname{BHMM}(n)$, or equivalently the classical closure of $\phi_n(\Theta_{\mathbb{C}})$, in $\mathbb{P}_p^{2^n-1}$.
- $I_{BHMM(n)}$, the Binary Hidden Markov ideal on n nodes, is the set of homogeneous polynomials which vanish on BHMM(n), i.e., the homogeneous defining ideal of $\overline{BHMM}(n)$. Elements of $I_{BHMM(n)}$ are called invariants of the model.

In summary, probability distributions arise from processes according to modeling maps, models are families of distributions arising from processes of a certain type, and the Zariski closure of each model is a variety whose geometry reflects the algebraic properties of the model. The ideal of the model is the same as the ideal of the variety: the definition of Zariski closure is the largest set which has the same ideal of vanishing polynomials as the model. In a rigorous sense (namely, the anti-equivalence of the categories of affine schemes

and rings), the variety encodes information about the "purely algebraic" properties of the model, i.e. properties that can be stated by the vanishing of polynomials.

The number of polynomials that vanish on any given set is infinite, but by the Hilbert Basis theorem, one can always find finitely many polynomials whose vanishing implies the vanishing of all the others. This is called a *generating set* for the ideal. To compute a generating set for $I_{BHMM(n)}$, we will need the following proposition:

Proposition 1. The ideal $I_{BHMM(n)}$ is the homogenization of $\ker(\phi_n^\# \circ \iota_n^\#)$ with respect to $p_{\Sigma} := \sum_{|v|=n} p_v$

Proof. The affine ideal $\ker(\phi_n^\# \circ \iota_n^\#)$ cuts out the Zariski closure X of $\iota_n \circ \phi_n(\Theta_{\mathbb{C}})$ in $\mathbb{C}_p^{2^n}$, and this closure lies in the hyperplane $\{p_{\Sigma}=1\}=\mathbb{C}_p^{2^n-1}$. Let X' be the projective closure of X in $\mathbb{P}_p^{2^n-1}$, so that I(X') is the homogenization of $\ker(\phi_n^\# \circ \iota_n^\#)$ with respect to p_{Σ} .

The cube Θ_{st} is Zariski dense in $\Theta_{\mathbb{C}}$, so $\iota_n \circ \phi_n(\Theta_{st})$ is Zariski dense in $\iota_n \circ \phi_n(\Theta_{\mathbb{C}})$, which is Zariski dense in X, which is Zariski dense in X'. Therefore $X' = \overline{\mathrm{BHMM}}(n)$, and $I(X') = I_{\mathrm{BHMM}(n)}$, as required.

2.3. HMMs with more visible states via BHMM(n)

All the results of this paper apply to HMMs with more than two visible states, using the following trick. Consider HMM(2, k, n), an HMM with 2 hidden states, k visible states $\alpha_1 \dots \alpha_k$, and n (consecutive) visible nodes. Such a hidden Markov process can be specified by a $2 \times k$ matrix E of emission probabilities, along with a 1×2 matrix π and a 2×2 matrix E describing the two-state hidden Markov chain as in (9). For each $\ell \in \{1 \dots, k\}$, we have a way to interpret this process as a BHM process by letting $\alpha_j = 1$ and $\alpha_i = 0$ for $i \neq j$. The resulting binary emission matrix is

$$E'(\ell) = \begin{bmatrix} 1 - E_{0\ell} & E_{0\ell} \\ 1 - E_{1\ell} & E_{1\ell} \end{bmatrix},$$

so as ℓ varies, we obtain all the entries E_{ij} as entries of an $E'(\ell)$. We shall remark throughout when results can be generalized to HMM(2, k, n) using this trick.

3. Defining equations of $\overline{BHMM}(3)$ and $\overline{BHMM}(4)$

Theorem 1. The homogeneous ideal $I_{BHMM(4)}$ of the binary hidden Markov variety $\overline{BHMM}(4)$ is minimally generated by 21 homogeneous quadrics and 29 homogeneous cubics.

Since Schönhuth [17] found numerically that his Hankel minors did not cut out BHMM(4) even set-theoretically, these equations are genuinely new invariants of the model. Moreover, they are not only applicable to BHMM(4), because a BHM process of length n > 4 can be marginalized to any 4 consecutive hidden-visible node pairs to obtain a BHM process of length 4. Thus, we have n - 3 linear maps from BHMM(n) to BHMM(4), each

of which allows us to write 21 quadrics and 29 cubics which vanish on BHMM(n). Finally, using Section 2.3, we can even obtain invariants of HMM(2, k, n) via the k different reductions to BHMM(n).

Our fastest derivation of Theorem 1 in Macaulay2 [11] uses the birational parametrization of Section 4, but in only a single step, so we defer the lengthier discussion of the parametrization until then. Modulo this dependency, the proof is described in Section 3.3, using moment coordinates (Section 3.1) and cumulant coordinates (Section 3.2).

In probability coordinates, the generators found for I_{BHMM(4)} had the following sizes:

- Quadrics $g_{2,1}, \ldots, g_{2,21}$: respectively 8, 8, 12, 14, 16, 21, 24, 24, 26, 26, 28, 32, 32, 41, 42, 43, 43, 44, 45, 72, 72 probability terms.
- Cubics $g_{3,1}, \ldots, g_{3,29}$: respectively 32, 43, 44, 44, 44, 52, 52, 56, 56, 61, 69, 71, 74, 76, 78, 81, 99, 104, 109, 119, 128, 132, 148, 157, 176, 207, 224, 236, 429 probability terms.

As a motivation for introducing moment coordinates, we note here that these generators have considerably fewer terms when written in terms of moments:

- Quadrics $g_{2,1}, \ldots, g_{2,21}$: respectively 4, 4, 4, 4, 6, 6, 6, 6, 6, 6, 6, 6, 6, 8, 8, 8, 8, 10, 10, 10, 17 moment terms.
- Cubics $g_{3,1}, \ldots, g_{3,29}$: respectively 5, 6, 6, 6, 6, 6, 6, 6, 6, 6, 8, 8, 8, 8, 10, 10, 10, 10, 10, 12, 12, 13, 14, 16, 18, 21, 27, 35 moment terms.

To give a sense of how these polynomials look in moment coordinates, the shortest quadric and cubic are

- $g_{2,1} = m_{23}m_{13} m_2m_{134} m_{13}m_{12} + m_1m_{124}$, and
- $g_{3,1} = m_{12}^3 2m_1m_{12}m_{123} + m_{\emptyset}m_{123}^2 + m_1^2m_{1234} m_{\emptyset}m_{12}m_{1234}.$

Let us compare this ideal with $I_{BHMM(3)}$, the homogeneous defining ideal of $\overline{BHMM}(3)$. Schönhuth [17] found that $I_{BHMM(3)}$ is precisely the ideal of 3×3 minors of the following matrix:

$$A_{3,3} = \begin{bmatrix} p_{000} + p_{001} & p_{000} & p_{100} \\ p_{010} + p_{011} & p_{001} & p_{101} \\ p_{100} + p_{101} & p_{010} & p_{110} \\ p_{110} + p_{111} & p_{011} & p_{111} \end{bmatrix}$$

$$(2)$$

Schönhuth defines an analogous matrix $A_{n,3}$ for $\overline{\text{BHMM}}(n)$, but then remarks that J. Hauenstein has found, using numerical rank deficiency testing [3] with the algebraic geometry package Bertini [2], that minors₃($A_{n,3}$) does not cut out $\overline{\text{BHMM}}(n)$ when n=4. In general, Schönhuth shows that $I_{\text{BHMM}(n)}=(\min s_3(A_{n,3}): \min s_2(B_{n,2}))$ for a particular 2×3 matrix $B_{n,2}$, but computing generators for this colon ideal is a costly operation, and so no generating set for $I_{\text{BHMM}(n)}$ was not found for any $n\geq 4$ by this method. Instead, here we will make use of moment coordinates and cumulant coordinates as exposited in [21].

3.1. Moment coordinates

Moments are particular linear expressions in probabilities. They can be derived from a moment generating function as in [21], but in our case, moments can be expressed simply by the following rule: we order $\{0,1\}^n$ by strict dominance, i.e. $v \ge w$ iff $v_i \ge w_i$ for all i, and then

$$m_v := \sum_{w \ge v} p_w \in R_{p,n} \tag{3}$$

Since all our variables are binary, with the usual algebraic statistical convention that a "+" subscript denotes an index to be summed over, we can view the conversion from moments to probabilities as "replacing zeros by + signs". For example, $m_{10010} = p_{1++1+}$. The ring elements $m_v \in R_{p,n}$ provide alternative linear coordinates on $\mathbb{P}_p^{2^n-1}$ in which it turns out that some previously intractable BHM computations are simplified and become feasible.

For a more compact notation, a binary string v of length n is the indicator function of a unique subset I of $[n] = \{1, \ldots, n\}$, so we also write m_I to represent m_v . For example, $m_{0000} = m_{\emptyset}$, $m_{1000} = m_1$, and $m_{0101} = m_{24}$. From (3) we can see that m_I actually represents a marginal probability: $m_I = \Pr(V_i = 1 \text{ for all } i \in I)$. Thus, in the context of BHMMs, no confusion results if we write m_I without specifying the value of n. To be precise, if $I \subseteq [n]$ and I' denotes I considered as a subset of [n'] for some n' > n, then

$$\phi_n^{\#}(m_I) = \phi_{n'}^{\#}(m_{I'}) \tag{4}$$

This can be seen in many ways, for example using the Baum formula for moments (Proposition 7) as explained in Section 5.3.

Just as for probabilities, for moments we define rings and spaces

$$R_{m,n} := \mathbb{C}[m_I \mid I \subseteq [n]] \qquad \qquad \mathbb{C}_m^{2^n} := \operatorname{Spec}(R_{m,n})$$

$$\overline{R}_{m,n} := R_{m,n} / \langle 1 - m_{\emptyset} \rangle \qquad \qquad \mathbb{C}_m^{2^n - 1} := \operatorname{Spec}(\overline{R}_{m,n}) \qquad (5)$$

$$\mathbb{P}_m^{2^n - 1} := \operatorname{Proj}(R_{m,n}),$$

To avoid having notation for too many ring isomorphisms, we adopt:

Convention 3.1. Using (3), we will usually treat m_I as a literal element of $R_{p,n}$, thus creating literal identifications

$$R_{m,n} = R_{p,n}$$

$$\overline{R}_{m,n} = \overline{R}_{p,n}$$

$$\mathbb{C}_m^{2^n} = \mathbb{C}_p^{2^n}$$

$$\mathbb{C}_m^{2^{n-1}} := \mathbb{C}_p^{2^{n-1}}$$

$$\mathbb{P}_m^{2^{n-1}} := \mathbb{P}_p^{2^{n-1}},$$
(6)

Note that, for example, we obtain natural ring inclusions

$$R_{m,n} \subseteq R_{m,n'}$$

whenever n < n', which respect the BHM maps ϕ_n because of (4).

As a first application of moment coordinates, we have

Proposition 2. The homogeneous ideal $I_{BHMM(3)}$ is generated in moment coordinates by the 3×3 minors of the matrix

$$A'_{3,3} = \begin{bmatrix} m_{000} & m_{000} & m_{100} \\ m_{010} & m_{001} & m_{101} \\ m_{100} & m_{010} & m_{110} \\ m_{110} & m_{011} & m_{111} \end{bmatrix} = \begin{bmatrix} m_{\emptyset} & m_{\emptyset} & m_{1} \\ m_{2} & m_{3} & m_{13} \\ m_{1} & m_{2} & m_{12} \\ m_{12} & m_{23} & m_{123} \end{bmatrix}$$

In particular, the projective variety $\overline{\rm BHMM}(3)$ is cut out by these minors.

Proof. Observe that Schönhuth's matrix $A_{3,3}$ in (2) is equivalent under elementary row/column operations to $A'_{3,3}$, so minors₃ $A'_{3,3} = \text{minors}_3 A_{3,3} = \text{I}_{\text{BHMM}(3)}$.

Proposition 3. The ideal $I_{BHMM(n)}$ is the homogenization of $\ker(\phi_n^{\#})$ with respect to m_{\emptyset} .

Proof. From Proposition 1 we know that $I_{BHMM(n)}$ is the homogenization of $\ker(\phi_n^\# \circ \iota_n^\#)$ with respect to $m_\emptyset = \sum_{|v|=n} p_v$. From (6), we can identify $\overline{R}_{m,4}$ with the polynomial subring of $R_{m,4}$ obtained by omitting m_\emptyset , so that $\ker(\phi_4^\# \circ \iota_4^\#) = \ker(\phi_4^\#) + \langle 1 - m_\emptyset \rangle$. Since the additional generator $1 - m_\emptyset$ homogenizes to 0, $\ker(\phi_4^\#)$ has the same homogenization as $\ker(\phi_4^\# \circ \iota_4^\#)$, hence the result.

3.2. Cumulant coordinates

Cumulants are non-linear expressions in moments or probabilities which seem to allow even faster computations with binary hidden Markov models. Let

$$R_{k,n} := \mathbb{C}[k_I \mid I \subseteq [n]]$$

$$\overline{R}_{k,n} := R_{k,n} / \langle k_{\emptyset} \rangle$$

$$\mathbb{C}_k^{2^n - 1} := \operatorname{Spec}(\overline{R}_{k,n})$$

where, as with moments, we may freely alternate between writing k_v and writing k_I , where I is the set of positions where 1 occurs in v. For building generating functions, let x_1, \ldots, x_n be indeterminates, and write $x^v = x^I$ for $x_1^{v_1} \cdots x_n^{v_n} = \prod_{i \in I} x_i$. Let J be the ideal generated by all the squares x_i^2 . Following [21], we define the moment and cumulant generating functions, respectively, as

$$f_m(x) := \sum_{I \subseteq [n]} m_I x^I \in \overline{R}_{m,n}[x]/J$$
 $f_k(x) := \sum_{I \subseteq [n]} k_I x^I \in \overline{R}_{k,n}[x]/J$

We now define changes of coordinates

$$\kappa_n: \mathbb{C}_m^{2^n-1} \to \mathbb{C}_k^{2^n-1} \qquad \qquad \kappa_n^{-1}: \mathbb{C}_k^{2^n-1} \to \mathbb{C}_m^{2^n-1}$$

by the formulae

$$\kappa_n^{\#}(f_k) = \log(f_m) = \frac{(f_m - 1)}{1} + \dots + (-1)^{n+1} \frac{(f_m - 1)^n}{n}$$

$$\kappa_n^{-\#}(f_m) = \exp(f_k) = 1 + \frac{(f_k)}{1} + \dots + \frac{(f_k)^n}{n!}$$
(7)

That is, we let $\kappa_n^{\#}(k_I)$ be the coefficient of x^I in the Taylor expansion of $\log f_m$ about 1, and let $\kappa_n^{\#}(m_I)$ be the coefficient of x^I in the Taylor expansion of $\exp f_k$ about 0. Note that in the relevant coordinate rings $\overline{R}_{m,n}$ and $\overline{R}_{k,n}$, $m_{\emptyset} = 1$ and $k_{\emptyset} = 0$. This is why we only need to compute the first n terms of each Talyor expansion: the higher terms all vanish modulo the ideal J.

Proposition 4. The expressions $\kappa_n^{\#}(k_I)$ and $\kappa_n^{-\#}(m_I)$, i.e. writing of cumulants in terms of moments and conversely, do not depend on n.

Proof. In [21], these formulae are re-expressed using Möbius functions, which do not depend on the generating functions above, and in particular do not depend on n.

3.3. Deriving I_{BHMM(4)} in Macaulay2

This section describes the proof of Theorem 1 using Macaulay2. These computations were carried out on a Toshiba Satellite P500 laptop running Ubuntu 10.04, with an Intel Core i7 Q740 .73 GHz CPU and 8gb of RAM. In light of Proposition 3, we will aim to compute $\ker(\phi_4^\# \circ \iota_4^\#)$, which can be understood geometrically as the (non-homogeneous) ideal of the standard affine patch of $\overline{\text{BHMM}}(4)$ where $m_\emptyset = \sum_{|v|=4} p_v = 1$. To reduce the number of variables, as in Proposition 3 we continue to make the identification

$$\overline{R}_{m,4} = \mathbb{C}[m_I | \emptyset \neq I \subseteq [4]] \subseteq R_{m,4}$$

We begin by providing Macaulay2 with the map $\phi_4^{\#}: \overline{R}_{m,4} \to \mathbb{C}[\theta]$ in moment coordinates (Section 3.1), because probability coordinates result in longer, higher degree expressions. This can be done by composing the expression of $\phi_n^{\#}(p_v)$ in Definition 3 with the expression of $m_v = m_I$ in (3), or alternatively using the Baum formula for moments (Proposition 7), which involves many fewer arithmetic operations.

Macaulay2 runs out of memory (8gb) trying to compute $\ker(\phi_4^{\#})$, and as expected, this memory runs out even sooner in probability coordinates, so we use cumulant coordinates instead (Section 3.2). We input

$$\kappa_4^\# : \overline{R}_{k,4} \to \overline{R}_{m,4}$$

using coefficient extraction from (7), and compute the composition $\phi_4^{\#} \circ \kappa_4^{\#}$. Then, it is possible to compute

$$\mathbf{I}_{k,4} := \ker(\phi_4^\# \circ \kappa_4^\#)$$

which takes **around 1.5 hours**. Alternatively, we can compute $I_{k,4}$ using the birational parameterization ψ_4 of Section 4 in place of ϕ_4 , which takes **less than 1 second** and yields 100 generators for $I_{k,4}$.

Subsequent computations run out of memory with this set of 100 generators, so we must take some steps to simplify it. Macaulay2's **trim** command reduces the number of generators of $\mathbf{I}_{k,4}$ to 46 in **under 1 second**. We then order these 46 generators lexicographically, first by degree and then by number of terms, and eliminate redundant generators in reverse order, which takes **19 seconds**. The result is an inclusion-minimal, non-homogeneous generating set for $\mathbf{I}_{k,4}$ with 35 generators: 24 quadrics and 11 cubics.

Now we compute $\mathbf{I}_{m,4} := \kappa^{\#}(I_{k,4}) = \kappa^{\#}(\ker(\phi_4^{\#} \circ \kappa_4^{\#})) = \ker(\phi_4^{\#})$, i.e., we push forward the 35 generators for $\mathbf{I}_{k,4}$ under the non-linear ring isomorphism $\kappa_4^{\#}$ to obtain 35 generators for $\mathbf{I}_{m,4} = \ker(\phi_4^{\#})$: 2 quadrics, 7 cubics, 16 quartics, 5 quintics, and 5 sextics. In **under 1 second**, Macaulay2's **trim** command computes a new set of 39 generators for $\mathbf{I}_{m,4}$ with lower degrees: 21 quadrics, 14 cubics, and 4 quartics, which turns out to **save around 1 hour** of computing time in what follows. These generators have many terms each, and eliminating redundant generators as in the previous paragraph turns out to be too slow to be worth it here, taking more than 2 hours, so we omit this step.

Finally, we apply Proposition 3 to compute $I_{BHMM(4)}$ as the homogenization of $I_{m,4}$ with respect to m_{\emptyset} . In Macaulay2, this is achieved by homogenizing the 39 generators for $I_{m,4}$ with respect to m_{\emptyset} and then saturating the ideal they generate with respect to m_{\emptyset} . This saturation operation takes about **29 minutes**, and yields a minimal generating set of 50 polynomials: 21 quadrics and 29 cubics. Since probabilities are linear in moments, their degrees are the same in probability coordinates. Moreover, since these are homogeneous generators for a homogeneous ideal, they are minimal in a very strong sense:

Corollary 1. Any inclusion-minimal homogeneous generating set for $I_{BHMM(4)}$ in probability or moment coordinates must contain exactly 21 quadrics and 29 cubics.

We still do not know a generating set for $I_{BHMM(5)}$. Macaulay2 runs out of memory (8gb) attempting to compute $I_{k,5}$, even using the birational parametrization of Section 4. The author has also attempted this computation using the *tree cumulants* of Smith and Zwiernik [19] in place of cumulants, but again Macaulay2 runs out of memory trying to compute the first kernel. Presumably the subsequent saturation step would be even more computationally difficult.

4. Birational parametrization of BHMMs

Theorem 2 (Birational Parameter Theorem). There is a generically two-to-one, dominant morphism $\Theta_{\mathbb{C}} \to \mathbb{C}^5$ such that, for each $n \geq 3$, the binary hidden Markov map ϕ_n factors uniquely as follows, and each $\psi_n : \mathbb{C}^5 \to \overline{\mathrm{BHMM}}(n)$ has a birational inverse map ρ_n :

$$\Theta_{\mathbb{C}} \xrightarrow{\psi_n} \mathbb{C}_p^{2^n - 1} \qquad \qquad \psi_n$$

$$\mathbb{C}^5 \xrightarrow{\overline{BHMM}(n)}$$

$$\rho_n$$

In particular, $\overline{BHMM}(n)$ is always a rational projective variety of dimension 5, i.e., birationally equivalent to \mathbb{P}^5 .

Using the reduction of Section 2.3, the same is true if we allow k > 2 visible states in the model and replace 5 by 3 + k. This theorem will be proven in Section 5.6 using trace algebras and the Baum formula for moments. In the course of this section and Section 5 we will exhibit formulae for ψ_n and their inverses ρ_n . The inverse map ρ_3 has a number of practical uses, to be explored in Section 6.

Our first step toward Theorem 2 is to re-parametrize $\Theta_{\mathbb{C}}$.

4.1. A linear reparametrization of $\Theta_{\mathbb{C}}$

Since the hidden variables H_t are never observed, there is no change in the final expression of p_v in Definition 3 if we swap the labels $\{0,1\}$ of all the H_t simultaneously. This swapping is equivalent to an action of the elementary permutation matrix $\sigma = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}$:

$$\mathbf{sw}: \Theta_{\mathbb{C}} \to \Theta_{\mathbb{C}}$$

$$\theta = (\pi, T, E) \mapsto (\pi \sigma, \, \sigma^{-1} T \sigma, \, \sigma^{-1} E)$$
(8)

(In our case $\sigma^{-1} = \sigma$, but the form above generalizes to permutations of larger hidden alphabets.) Hence we have that $\Pr(v \mid \pi, T, E) = \Pr(v \mid \mathbf{sw}(\pi, T, E))$, i.e. $\phi_n = \phi_n \circ \mathbf{sw}$.

We will make essential use of a linear parametrization of $\Theta_{\mathbb{C}}$ in which **sw** has a simple form. Our new parameters will be $\eta_0 := (a_0, b, c_0, u, v_0)$, with subscript 0's to be explained shortly. Although we have already used the letter v at times to represent visible binary strings, we hope that the context will be clear enough to avoid confusion between these usages. We let

$$\pi = \frac{1}{2} \begin{bmatrix} 1 - a_0, & 1 + a_0 \end{bmatrix}$$

$$T = \frac{1}{2} \begin{bmatrix} 1 + b - c_0, & 1 - b + c_0 \\ 1 - b - c_0, & 1 + b + c_0 \end{bmatrix} \qquad E = \begin{bmatrix} 1 - u + v_0, & u - v_0 \\ 1 - u - v_0, & u + v_0 \end{bmatrix}$$
(9)

(The rightmost column of E is made intentionally homogeneous in the new parameters.) We can linearly solve for η_0 in terms of θ by $a_0 = \pi_1 - \pi_0$ etc., so in fact (a_0, b, c_0, u, v_0) generate the parameter ring $\mathbb{C}[\theta]$. In these coordinates, sw acts by

$$a_0 \mapsto -a_0,$$
 $b \mapsto b,$ $c_0 \mapsto -c_0,$ $u \mapsto u,$ $v_0 \mapsto -v_0$

In other words, swapping the signs of the subscripted variables a_0, c_0, v_0 has the same effect as acting on the matrices π, T, E by σ as in (8), i.e., relabeling the hidden alphabet.

4.2. Introducing the birational parameters

Since $\phi_n \circ \mathbf{sw} = \phi_n$, by classical invariant theory the ring map $\phi_n^\# : \overline{R}_{p,n} \to \mathbb{C}[\theta]$ must land in the subring of invariants $\mathbb{C}[\theta]^{\mathbf{sw}} = \mathbb{C}[b, u, a_0^2, c_0^2, v_0^2, a_0c_0, a_0v_0, c_0v_0]$. However, $\phi_n^\#$ in fact factors through a smaller subring, conveniently generated by 5 elements:

Lemma 1 (Parameter Subring Lemma). For all $n \geq 3$, the ring map $\phi_n^{\#}$ lands in the subring

$$\mathbb{C}[\eta] := \mathbb{C}[a,b,c,u,v]$$

of $\mathbb{C}[\theta]$, where $a = a_0 v_0$, $c = c_0 v_0$, $v = v_0^2$.

The proof of this key lemma will be given in Section 5.5 after introducing trace algebras. To interpret its geometric consequences, write $\mathfrak{q}^{\#}$ for the subring inclusion

$$\mathfrak{q}^{\#}: \mathbb{C}[\eta] \hookrightarrow \mathbb{C}[\theta]$$

$$a \mapsto a_0 v_0, \quad b \mapsto b, \quad c \mapsto c_0 v_0, \quad u \mapsto u, \quad v \mapsto v_0^2,$$

write $\psi_n^{\#}: \overline{R}_{p,n} \to \mathbb{C}[\eta]$ for the factorization of $\phi_n^{\#}$ through $\mathfrak{q}^{\#}$, and write $\Theta_{\mathbb{C}}' := \operatorname{Spec} \mathbb{C}[\eta]$, so $\Theta_{\mathbb{C}}' \simeq \mathbb{C}^5$. The result:

Corollary 2. The following diagram of dominant maps commutes

$$\Theta_{\mathbb{C}} \xrightarrow{\mathfrak{q}} \Theta'_{\mathbb{C}} \xrightarrow{\psi_n} \overline{\mathrm{BHMM}}(n)$$

$$\phi_n$$

and \mathfrak{q} is generically two-to-one.

This corollary in particular implies the first part of the Birational Parameter Theorem (2), by taking $\mathfrak{q}:\Theta_{\mathbb{C}}\to\Theta'_{\mathbb{C}}\simeq\mathbb{C}^5$ as the generically 2:1 map.

Remark 2. The map \mathfrak{q} is only dominant, and not surjective; for example, it misses the point (1,0,0,0,0).

Corollary 3. For all $n \geq 3$, $\overline{BHMM}(n) = \overline{image}(\overline{\iota}_n \psi_n)$.

Proof. Since
$$\mathfrak{q}$$
 is dominant, $\overline{\mathrm{image}}(\bar{\iota}_n\psi_n) = \overline{\mathrm{image}}(\bar{\iota}_n\psi_n\mathfrak{q}) = \overline{\mathrm{image}}(\bar{\iota}_n\phi_n) =: \overline{\mathrm{BHMM}}(n)$.

The unique factorization map $\psi_n^{\#}$ can be computed directly in Macaulay2 for small n. The expressions in moment coordinates are simpler than in probabilities, so we present these in the following proposition.

Proposition 5. The map $\psi_3^{\#}$ is given in moment coordinates by

$$m_{\emptyset} = m_{000} \mapsto 1$$

$$m_1 = m_{100} \mapsto a + u$$

$$m_{2} = m_{010} \mapsto ab + c + u$$

$$m_{3} = m_{001} \mapsto ab^{2} + bc + c + u$$

$$m_{12} = m_{110} \mapsto abu + ac + au + cu + u^{2} + bv$$

$$m_{13} = m_{101} \mapsto ab^{2}u + abc + bcu + b^{2}v + ac + au + cu + u^{2}$$

$$m_{23} = m_{011} \mapsto ab^{2}u + abc + abu + bcu + c^{2} + 2cu + u^{2} + bv$$

$$m_{123} = m_{111} \mapsto ab^{2}u^{2} + 2abcu + abu^{2} + bcu^{2} + b^{2}uv + ac^{2} + 2acu + c^{2}u + au^{2} + 2cu^{2} + u^{3} + abv + bcv + 2buv$$

We will eventually prove the Birational Parameter Theorem (2) by marginalization to the case n=3, which we can prove here:

Proposition 6. The following triangular set of equations hold on the graph of ψ_3 , after clearing denominators, and can thus be used to recover parameters from observed moments where the denominators are non-zero:

$$b = \frac{m_3 - m_2}{m_2 - m_1}$$

$$u = \frac{m_1 m_3 - m_2^2 + m_{23} - m_{12}}{2(m_3 - m_2)}$$

$$a = m_1 - u$$

$$c = a - ba + m_2 - m_1$$

$$v = a^2 - \frac{m_1 m_2 - m_{12}}{b}$$

(This proposition and the following corollary actually hold for all ϕ_n with $n \geq 3$, because of Proposition 8, and by Section 2.3, these same formulae can be used to recover parameters for HMM(2,k,n) when k>2 as well.)

Proof. These equations can be checked with direct substitution by hand from Proposition 5. Regarding the derivation, they can be obtained in Macaulay2 by computing two Gröbner bases of the elimination ideal $I = \langle m_v - \phi_3(m_v) | v \in \{0,1\}^3 \rangle$ over the ring $\mathbb{C}_m^{2^3}$, in Lex monomial order: once in the ring $R_{m,3}[v,c,a,b,u]$, and once in $R_{m,3}[v,c,u,b,a]$. Each variable occurs in the leading term of some generator in one of these two bases with a simple expression in moments as its leading coefficient. We solve each such generator (set to 0) for the desired parameter.

Corollary 4. The map $\psi_3: \mathbb{C}^5 \to \overline{BHMM}(3)$ has a birational inverse ρ_3 . The map $\rho_3^{\#}$ on moment coordinate functions is given by:

$$a \mapsto \frac{m_2^2 + m_3 m_1 - 2m_2 m_1 - m_{23} + m_{12}}{2(m_3 - m_2)} \qquad u \mapsto \frac{-m_2^2 + m_3 m_1 + m_{23} - m_{12}}{2(m_3 - m_2)}$$
$$b \mapsto \frac{m_3 - m_2}{m_2 - m_1} \qquad v \mapsto \frac{\text{num}(v)}{4(m_3 - m_2)^2}$$

$$c \mapsto \frac{\operatorname{num}(c)}{2(m_2 - m_1)(m_3 - m_2)}, \text{ where}$$

$$\begin{aligned} \mathrm{num}(c) &= -\,m_1 m_2^2 + m_1^2 m_3 + m_2^2 m_3 - m_1 m_3^2 - m_1 m_{12} \\ &\quad + 2 m_2 m_{12} - m_3 m_{12} + m_1 m_{23} - 2 m_2 m_{23} + m_3 m_{23}, \ and \\ \mathrm{num}(v) &= m_2^4 - 2 m_1 m_2^2 m_3 + m_1^2 m_3^2 - 2 m_2^2 m_{12} - 2 m_1 m_3 m_{12} + 4 m_2 m_3 m_{12} \\ &\quad + 4 m_1 m_2 m_{23} - 2 m_2^2 m_{23} - 2 m_1 m_3 m_{23} + m_{12}^2 - 2 m_{12} m_{23} + m_{23}^2. \end{aligned}$$

Proof. This can be derived by substituting the solutions for u, a, and b in the previous propositions into the subsequent solutions for a, c, and v. Alternatively, it can be checked by direct substitution in Macaulay2, i.e., one computes that $\psi_3^{\#} \circ \rho^{\#}(\theta) = \theta$ for each birational parameter $\theta \in \{a, b, c, u, v\}$.

The expressions in Corollary 4 are considerably simpler in moment coordinates than in probabilities. Comparing the number of terms, the numerators for a, b, c, u, v respectively have sizes 5, 2, 10, 4, and 12 in moment coordinates, versus sizes 22, 4, 56, 22, and 190 in probability coordinates. This explains in part why Macaulay2's Gröbner basis computations execute in moment coordinates with much less time and memory.

4.3. Statistical interpretation of the birational inverse ρ_3

It turns out that the factors appearing in the denominators of Corollary 4 defining ρ_3 have simple factorizations in terms of the rational and birational parameters:

• $m_3 - m_2$ appears in the denominator of all $\rho_3(\theta)$ except $\rho_3(b)$, and

$$m_3 - m_2 \stackrel{\psi_3}{\mapsto} (b)(ab - a + c) \stackrel{\mathfrak{q}}{\mapsto} (b)(v_0)(a_0b - a_0 + c_0)$$

• $m_2 - m_1$ appears in the denominator of $\rho_3(b)$ and $\rho_3(c)$, and

$$m_2 - m_1 \stackrel{\psi_3}{\mapsto} ab - a + c \stackrel{\mathfrak{q}}{\mapsto} (v_0)(a_0b - a_0 + c_0)$$

Let us pause to reflect on the meaning of these factors.

• The factor v_0 occurs in $det(E) = 2v_0$, hence $v = v_0^2 = 0$ iff the hidden Markov chain has "no effect" on the observed variables. The image locus $\phi_3(\{v_0 = 0\})$ can thus be modeled by a sequence of IID coin flips with distribution $E_0 = E_1 = (1 - u, u)$, so the BHMM is an unlikely model choice. This is a **one-dimensional submodel**, parametrizable by $u \in [0, 1]$, with a regular (everywhere-defined) inverse given simply by $u = m_1$. Denote this model by BIID(n).

- The factor b occurs in det(T) = b, hence b = 0 iff each hidden node has "no effect" on the subsequent hidden nodes. In this case, the observed process can be modeled as a sequence of independent coin flips, the first flip having distribution $(1 \alpha, \alpha) := \pi E$ and subsequent flips being IID having distribution $(1 \beta, \beta) := T_0E = T_1E$. The image locus $\phi_3(\{b = 0\})$ is hence a **two-dimensional submodel**, parametrizable by $(\alpha, \beta) \in [0, 1]^2$, with a regular inverse given by $\alpha = m_1, \beta = m_2$. Denote this model by BINID(n), for "binary independent nearly identically distributed" model, and note that BINID $(n) \supseteq BIID(n)$ by setting $\alpha = \beta$.
- The factor $a_0b a_0 + c_0$ occurs in $\pi T \pi = \frac{1}{2}(-a_0b + a_0 c_0, a_0b a_0 + c_0)$. Hence $a_0b - a_0 + c_0 = 0$ iff π is a fixed point of T, i.e. the hidden Markov chain is at equilibrium. We may define the Equilibrium Binary Hidden Markov model by restricting ϕ_n to the locus $\{a_0b - a_0 + c_0 = 0\}$), which turns out to yield a **four-dimensional submodel** for each $n \geq 3$. Denote this submodel by EBHMM(n).

It can be easily shown, with the same methods used here for BHMM(n), that EBHMM(n) itself has a birational parametrization by $(a_0v_0, b, u, v_0^2) = (a, b, u, v)$, where $a_0, b \in [-1, 1]$, $c_0 := a_0(1-b) \in [|b|-1, 1-|b|]$, $v_0 \in [0, 1]$, and $u \in [|v_0|, 1-|v_0|]$, with an inverse parametrization given by

$$b = \frac{m_1^2 - m_{13}}{m_1^2 - m_{12}} \qquad u = \frac{2m_1 m_{12} - m_1 m_{13} - m_{123}}{2(m_1^2 - m_{13})}$$

$$a = m_1 - u \qquad v = \frac{a^2 b - m_1^2 + m_{12}}{b}$$

The newly occurring denominators here are $m_1^2 - m_{12} = (b)(a^2 - v) = (b)(v_0)^2(a_0^2 - 1)$ and $m_1^2 - m_{13} = (b)^2(a^2 - v) = (b)(v_0)^2(a_0^2 - 1)$. It easy to check that the only points of EBHMM(n) where these expressions vanish are points that lie in BINID(n). Thus, for $n \geq 3$, BHMM(n) can be stratified as a union of three statistically meaningful submodels

each of which has an everywhere-defined inverse parametrization.

4.4. Computational advantages of moments, cumulants, and birational parameters

Our approach has been to work with moments m_v and cumulants k_v instead of probabilities p_v , and the birational parameters a, b, c, u, v instead of the matrix entries π_1, t_{i1}, e_{i1} . Other than the theoretical advantage that the model map is generically injective on the birational parameter space, significant computation gains in Macaulay2 also result from these choices (see Section 3.3 for laptop specifications):

- Computing $\ker \psi_3 = \ker \phi_3$, the affine defining ideal of $\overline{\text{BHMM}}(3)$, took less than 1 second in Macaulay2 when using the birational parameters, compared to 25 seconds when using the matrix entries and moments, and 15 minutes when using the matrix entries and probabilities.
- Computing $\ker \psi_4 = \ker \phi_4$, the affine defining ideal of $\overline{BHMM}(4)$ took less than 1 second in Macaulay2 when using the birational parameters and *cumulant coordinates* [21], compared to 1.5 hours when using the matrix entries and cumulant coordinates, and running out of memory (8gb) when using the matrix entries and probabilities.

5. Parametrizing BHMMs though a trace variety

In this section, we exhibit a parametrization of every BHMM through a particular trace variety called Spec $C_{2,3}$, which itself can be embedded in \mathbb{C}^{10} . We use these coordinates to prove the Birational Parameter Theorem (2) and the Parameter Subring Lemma (1), which were stated without proof.

For this, we will define a map ϕ_{∞} through which all the ϕ_n factor, and using a version of the Baum formula for moments, we factor this map further though Spec $C_{2,3}$. Then we use a finite set 10 of generators of the ring $C_{2,3}$ exhibited by [18] to show that the image of ϕ_{∞} lands in the desired subring $\mathbb{C}[\eta]$, and write ψ_{∞} for the factorization. Finally, by marginalizing to the case n=3, we obtain a birational inverse for ψ_n from the map ρ_3 given in Corollary 4.

5.1. Marginalization maps

For each pair of integers $n' \geq n \geq 1$, the marginalization map $\mu_n^{n'}: \mathbb{C}_p^{2^{n'}} \to \mathbb{C}_p^{2^n}$ is given by

$$\mu_n^{n'^{\#}}(p_v) := \sum_{|w|=n'-n} p_{vw}$$

These restrict to maps $\mu_n^{n'}: \mathbb{C}_p^{2^{n'}-1} \to \mathbb{C}_p^{2^n-1}$, and define rational maps $\mu_n^{n'}: \mathbb{P}_p^{2^{n'}-1} \dashrightarrow \mathbb{P}_p^{2^n-1}$. In moment coordinates, these maps are actually coordinate projections: $\mu_n^{n'\#}(m_v) = m_{v\bar{0}}$ where $\bar{0}$ denotes a sequence of n'-n zeros. In fact, using the subset notation for moments m_I , the corresponding ring maps are literal inclusions: $\mu_n^{n'\#}(m_I) = m_I$. In other words, $\mu_n^{n'}: \mathbb{C}_m^{2^{n'}} \to \mathbb{C}_m^{2^n}$ is just the map which forgets those m_I where $I \nsubseteq [n]$.

5.2. The Baum formula for moments

Equation (1) involves $O(2^n)$ addition operations. There is a faster way to compute $\phi_n^{\#}(p_v)$, using O(n) arithmetic operations, by treating the BHM process as a *finitary* process [17]. We define two new matrices[†]

$$(P_i)_{jk} := E_{ji}T_{jk} = \Pr(V_t = i \text{ and } H_{t+1} = k \mid H_t = j \text{ and } \pi, E, T), \text{ that is,}$$

 $^{^{\}dagger}P$ can be thought of naturally as a $2\times2\times2$ tensor, but we will not make use of this interpretation.

$$P_0 := \begin{bmatrix} T_{00}E_{00} & T_{01}E_{00} \\ T_{10}E_{10} & T_{11}E_{10} \end{bmatrix} \quad \text{and} \quad P_1 := \begin{bmatrix} T_{00}E_{01} & T_{01}E_{01} \\ T_{10}E_{11} & T_{11}E_{11} \end{bmatrix}$$

Writing $\mathbb{1}$ for the vector $\begin{pmatrix} 1 \\ 1 \end{pmatrix}$ we obtain the matrix expression $\phi^{\#}(p_v) = \pi P_{v_1} P_{v_2} \cdots P_{v_v} \mathbb{1}$ which involves only 4n+2 multiplications and 2n+1 additions. This is known as the Baum formula. We can rewrite this formula as a trace product of 2×2 matrices:

$$\phi^{\#}(p_v) = \operatorname{trace}(\pi P_{v_1} P_{v_2} \cdots P_{v_n} \mathbb{1}) = \operatorname{trace}((\mathbb{1}\pi) P_{v_1} P_{v_2} \cdots P_{v_n})$$

To create an analogue of this formula in moment coordinates, we let

$$M_0 := P_0 + P_1 = T$$
 $M_1 := P_1$ $M_2 := \mathbb{1}\pi = \begin{bmatrix} \pi_0 & \pi_1 \\ \pi_0 & \pi_1 \end{bmatrix}$

Proposition 7 (Baum formula for moments). The binary hidden Markov map ϕ_n can be written in moment coordinates as

$$\phi_n^{\#}(m_v) = \operatorname{trace}(M_2 M_{v_1} M_{v_2} \cdots M_{v_n})$$

For example, $\phi_n^{\#}(m_{01001}) = \operatorname{trace}(M_2 M_0 M_1 M_0 M_0 M_1).$

Proof. By our definition of m_v (3), we have

$$\phi_n^{\#}(m_v) = \sum_{w \ge v} \phi_n^{\#}(p_w) = \sum_{w \ge v} \operatorname{trace}((\mathbb{1}\pi) P_{w_1} P_{w_2} \cdots P_{w_n})$$

$$= \operatorname{trace}\left((\mathbb{1}\pi) \left(\sum_{w_1 \ge v_1} P_{w_1}\right) \left(\sum_{w_2 \ge v_2} P_{w_2}\right) \cdots \left(\sum_{w_n \ge v_n} P_{w_n}\right)\right)$$

$$= \operatorname{trace}(M_2 M_{v_1} M_{v_2} \cdots M_{v_n})$$

5.3. Truncation and ϕ_{∞}

Proposition 8. The binary hidden Markov maps ϕ_n form a directed system of maps under marginalization, meaning that, for each $n' \geq n \geq 1$, the following diagrams commute:

$$\Theta_{\mathbb{C}} \bigvee_{\phi_{n}}^{\phi_{n'}} \bigvee_{\mathbb{C}_{m}^{2^{n'}-1}}^{\mathbb{C}_{m}^{2^{n'}-1}} \qquad \qquad \bigoplus_{\phi_{n'}}^{\phi_{n'}} \overline{R}_{m,n'} \\ \bigoplus_{\mathbb{C}_{m}^{2^{n}-1}}^{\phi_{n'}^{\#}} \bigvee_{\phi_{n}^{\#}} \overline{R}_{m,n}$$

Proof. This can be seen directly from the definition of ϕ_n using (1) and of m_v in (3). Alternatively, observe that because $M_0 = T$ is stochastic, $M_0M_2 = M_2$, so for any sequence $\overline{0}$ of length n' - n, the Baum formula for moments (Proposition 7) implies that

$$\phi_{n'}^{\#}(m_{v\overline{0}}) = \phi_n^{\#}(m_v) \tag{10}$$

Thus, to compute ϕ_n for all n, it is only necessary to compute those $\phi_n^\# m_{v'}$ where v' ends in 1. Motivated by this observation, let $\overline{R}_{m,\infty} := \mathbb{C}[m_{v1} \mid v \in \{0,1\}^n$ for some $n \geq 0$ $= \mathbb{C}[m_1, m_{01}, m_{11}, m_{001}, m_{101}, m_{011}, \ldots]$, which in subset index notation is simply

$$\overline{R}_{m,\infty} := \mathbb{C}[m_I \mid I \subseteq [n] \text{ for some } n \ge 0]$$
$$= \mathbb{C}[m_1, m_2, m_{12}, m_3, m_{13}, m_{23}, \ldots]$$

Then we define $\phi_{\infty}: \Theta_{\mathbb{C}} \to \operatorname{Spec} \overline{R}_{m,\infty}$ and $\phi_{\infty}^{\#}: \mathbb{C}[\theta] \leftarrow \overline{R}_{m,\infty}$ by the formula $\phi_{\infty}^{\#}(m_{v1\overline{0}}) := \phi_{\operatorname{length}(v1)}^{\#}(m_{v1})$, i.e.

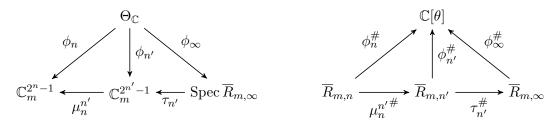
$$\phi_{\infty}^{\#}(m_I) := \phi_{\operatorname{size}(I)}^{\#}(m_I) \tag{11}$$

Note that by locating the position of the last 1 in a binary sequence $v' \neq 0...0$, we can write v' in the form $v1\overline{0}$ for a unique string v (possibly empty if v' = 1), so this map is well-defined. By the same principle, for each n we can also define a "truncation" map τ : Spec $\overline{R}_{m,\infty} \to \mathbb{C}_m^{2^n-1}$ by $\tau^{\#}(m_{v1\overline{0}}) := m_{v1}$, which, in subset index notation, is a literal ring inclusion:

$$\tau^{\#}(m_I) := m_I \tag{12}$$

With this definition, $\phi_n^{\#}$ factorizes as $\phi_n^{\#} = \phi_{\infty}^{\#} \circ \tau_n^{\#}$. We can summarize this and Proposition 8 as follows:

Proposition 9. For all $n' \ge n \ge 1$, the following diagrams commute:



Remark 3. These diagrams exhibit the rings $\overline{R}_{m,n}$ and maps $\phi_n^{\#}$ as a directed system under the inclusion maps ${\mu_n^{n'}}^{\#}$, such that $\overline{R}_{m,\infty} = \operatorname{colim}_{n\to\infty} \overline{R}_{m,n}$ and $\phi_\infty^{\#} = \lim_{n\to\infty} \phi_n^{\#}$.

Now, to prove that ϕ_n factors through \mathfrak{q} , we need only show that ϕ_{∞} does.

5.4. Factoring ϕ_{∞} through a trace variety

Let X_0, X_1, X_2 be 2×2 matrices of indeterminates,

$$X_0 = \begin{bmatrix} x_{000} & x_{001} \\ x_{010} & x_{011} \end{bmatrix} \qquad X_1 = \begin{bmatrix} x_{100} & x_{101} \\ x_{110} & x_{111} \end{bmatrix} \qquad X_2 = \begin{bmatrix} x_{200} & x_{201} \\ x_{210} & x_{211} \end{bmatrix}$$

and following the notation of [8], $\Omega_{2,3} := \mathbb{C}[\text{entries of } X_0, X_1, X_2]$ denotes the polynomial ring on the entries x_{ijk} of these three 2×2 matrices. The trace algebra $C_{2,3}$ is defined

as the subring of $\Omega_{2,3}$ generated by the traces of products of these matrices, $C_{2,3} := \mathbb{C}[\operatorname{trace}(X_{i_1}X_{i_2}\cdots X_{i_r}) \mid r \geq 1] \subseteq \Omega_{2,3}$ and we refer to Spec $C_{2,3}$ as a trace variety. We write

$$\nu: \operatorname{Spec} \Omega_{2,3} \to \operatorname{Spec} C_{2,3}$$
 and $\nu^{\#}: C_{2,3} \hookrightarrow \Omega_{2,3}$

for the natural dominant map and corresponding ring inclusion. To relate these varieties to binary HMMs, we define two new maps $\omega^{\#}: \Omega_{2,3} \to \mathbb{C}[\theta]$ and $\xi^{\#}: \overline{R}_{m,\infty} \to C_{2,3}$ by

$$\omega^{\#}(X_i) := M_i$$
 and $\xi^{\#}(m_{v1}) := \operatorname{trace}\left(\left(X_2 \prod_{i \in v} X_i\right) X_1\right).$

Proposition 10 (Baum factorization). The ring map $\phi_{\infty}^{\#}$ factorizes as $\phi_{\infty}^{\#} = \omega^{\#} \circ \nu^{\#} \circ \xi^{\#}$, i.e., the following diagram commutes:

$$\Theta_{\mathbb{C}} \xrightarrow{\phi_{\infty}} \operatorname{Spec} \overline{R}_{m,\infty}$$

$$\omega \qquad \qquad \qquad \downarrow \qquad \qquad \xi$$

$$\operatorname{Spec} \Omega_{2,3} \xrightarrow{\nu} \operatorname{Spec} C_{2,3}$$

Proof. This is just a restatement of the Baum formula for moments (Proposition 7):

$$\omega^{\#}(\nu^{\#}(\xi^{\#}(m_{v1}))) = \omega^{\#} \operatorname{trace}\left(X_{2} \prod_{i \in v1} X_{i}\right) = \operatorname{trace}\left(M_{2} \prod_{i \in v1} M_{i}\right)$$
$$= \phi_{length(v1)}^{\#}(m_{v1}) = \phi_{\infty}^{\#}(m_{v1})$$

5.5. Proving the Parameter Subring Lemma

We begin by seeking a factorization of the map $\omega^{\#} \circ \nu^{\#}$. For this we apply the following commutative algebra result of Sibirskii on the trace algebras $C_{2,r}$:

Proposition 11 (Sibirskii, 1968). The trace algebra $C_{2,r}$ is generated by the elements

$$trace(X_i) : 0 \le i \le r$$
$$trace(X_i X_j) : 0 \le i \le j \le r$$
$$trace(X_i X_j X_k) : 0 \le i < j < k \le r$$

Corollary 5. The algebra $C_{2,3}$ is generated by the 10 elements

$$\begin{aligned} &\operatorname{trace}(X_0), \operatorname{trace}(X_1), \operatorname{trace}(X_2), \\ &\operatorname{trace}(X_0^2), \operatorname{trace}(X_1^2), \operatorname{trace}(X_2^2), \operatorname{trace}(X_0X_1), \operatorname{trace}(X_0X_2), \operatorname{trace}(X_1X_2), \\ &\operatorname{trace}(X_0X_1X_2) \end{aligned}$$

Proposition 12. The ring map $\omega^{\#} \circ \nu^{\#}$ factors through the inclusion

$$\mathfrak{q}^{\#} : \mathbb{C}[\eta] := \mathbb{C}[a, b, c, u, v] \hookrightarrow \mathbb{C}[\theta] := \mathbb{C}[a_0, b, c_0, u, v_0],$$

i.e. we can write $\omega^{\#} \circ \nu^{\#} = \mathfrak{q}^{\#} \circ \mathfrak{r}^{\#}$ so that the following diagram commutes:

$$\begin{array}{ccc}
\Theta_{\mathbb{C}} & & \mathfrak{q} & & \Theta'_{\mathbb{C}} \\
\downarrow \omega & & & \downarrow \mathfrak{r} \\
\operatorname{Spec} \Omega_{2,3} & & \longrightarrow & \operatorname{Spec} C_{2,3}
\end{array}$$

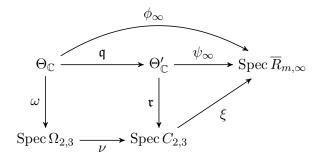
Proof. We apply $\omega^{\#}$ to the ten generators of $C_{2,3}$ given in Corollary 5 and check that they land in $\mathbb{C}[\eta]$. Explicit, we find that:

$$\begin{aligned} & \operatorname{trace}(M_0) = b + 1 & \operatorname{trace}(M_1) = bu + c + u & \operatorname{trace}(M_2) = 1 \\ & \operatorname{trace}(M_0^2) = b^2 + 1 & \operatorname{trace}(M_1^2) = b^2 u^2 + 2bcu + c^2 + 2cu + u^2 + 2bv \\ & \operatorname{trace}(M_2^2) = 1 & \operatorname{trace}(M_0 M_1) = b^2 u + bc + c + u & \operatorname{trace}(M_0 M_2) = 1 \\ & \operatorname{trace}(M_1 M_2) = a + u & \operatorname{trace}(M_0 M_1 M_2) = ab + c + u \end{aligned}$$

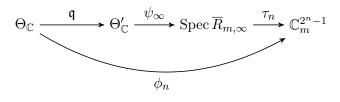
Now, by letting $\psi_{\infty}^{\#} := \mathfrak{r}^{\#} \circ \xi^{\#}$ we may factor the ring map $\phi_{\infty}^{\#}$ as

$$\phi_{\infty}^{\#} = \omega^{\#} \circ \nu^{\#} \circ \xi^{\#} = \mathfrak{q}^{\#} \circ \mathfrak{r}^{\#} \circ \xi^{\#} = \mathfrak{q}^{\#} \circ \psi_{\infty}^{\#}.$$

Corollary 6. The following diagram commutes:



Proof. [Proof of the Parameter Subring Lemma] Proposition 9 and Corollary 6 together imply that the following diagrams commute:



$$\mathbb{C}[\theta] \xleftarrow{\mathfrak{q}^{\#}} \mathbb{C}[\eta] \xleftarrow{\psi_{\infty}^{\#}} \overline{R}_{m,\infty} \xleftarrow{\tau_{n}^{\#}} \overline{R}_{m,n}$$

$$\phi_{n}^{\#}$$

In particular, the map $\phi_n^{\#}$ factors through $\mathbb{C}[\eta]$, as required.

5.6. Proving the Birational Parameter Theorem (2)

Recall that Corollary 2 implies the first part of the Birational Parameter Theorem (2), by taking

$$\mathfrak{q}:\Theta_{\mathbb{C}}\longrightarrow\Theta'_{\mathbb{C}}$$

as the generically 2:1 map. Thus, it remains to show that the maps

$$\psi_n:\Theta'_{\mathbb{C}}\longrightarrow \overline{\mathrm{BHMM}}(n)$$

have birational inverses ρ_n . The inverse map ρ_3 was already exhibited in Corollary 4, and we obtain ρ_n by marginalization: let

$$\rho_n = \rho_3 \circ \mu_3^n.$$

Let $U \subseteq \Theta'_{\mathbb{C}}$ be the Zariski open set on which ψ_3 is an isomorphism with inverse ρ_3 . Consider the set $\psi_n(U) \subseteq \overline{\mathrm{BHMM}}(n)$. It is Zariski dense in $\overline{\mathrm{BHMM}}(n)$, and by Chevalley's theorem (12, EGA IV, 1.8.4), it is constructible, so it must contain a dense open set $W' \subseteq \overline{\mathrm{BHMM}}(n)$. Now let $W = \psi_n^{-1}(W')$, so we have $\psi_n(W) = W' \subseteq \psi_n(U)$.

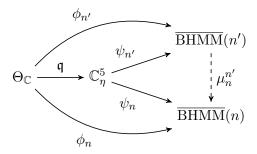
Proposition 13. $\rho_n \circ \psi_n = Id \text{ on } W \text{ and } \psi_n \circ \rho_n = Id \text{ on } W'.$

Proof. Suppose $\widehat{\eta} \in W$. Then $\rho_n \circ \psi_n(\widehat{\eta}) = \rho_3 \circ \mu_3^n \circ \psi_n(\widehat{\eta}) = \rho_3 \circ \psi_3(\widehat{\eta}) = \widehat{\eta}$ since $\widehat{\eta} \in U$. Now suppose $\widehat{p} \in W'$, so $\widehat{p} = \psi_n(\widehat{\eta})$ for some $\widehat{\eta} \in W$. Then, applying Proposition 8,

$$\psi_n \circ \rho_n(\widehat{p}) = \psi_n \circ \rho_n \circ \psi_n(\widehat{\eta}) = \psi_n \circ \rho_3 \circ \mu_3^n \circ \psi_n(\widehat{\eta})$$
$$= \psi_n \circ \rho_3 \circ \psi_3(\widehat{\eta}) = \psi_n(\widehat{\eta}) = \widehat{p}$$

This completes the proof of the Birational Parameter Theorem (2). In fact we have also proven the following:

Theorem 3. For any $n' \ge n \ge 3$, there is a commutative diagram of dominant maps:



6. Applications and future directions

Besides attempting to compute a set of generators for $I_{BHMM(5)}$, there are many other questions to be answered about HMMs that can be approached immediately with the techniques of this paper.

6.1. A nonnegative distribution in $\overline{BHMM}(3)$ but not BHMM(3)

It turns out that not all of the probability distributions (non-negative real points) of $\overline{\rm BHMM}(n)$ lie in the model BHMM(n). In other words, $\overline{\rm BHMM}(n) \cap \Delta_p^{2^n-1} \neq \rm BHMM(n)$, so the model must be cut out by some non-trivial inequalities inside the simplex. To illustrate this, the following real point $\widehat{\theta}$ of $\Theta_{\mathbb{C}}$ does not lie in $\Theta_{\rm st}$, but maps under ϕ_3 to a point \widehat{p} of Δ_p^7 :

$$\widehat{\theta} = (\widehat{\pi}, \widehat{T}, \widehat{E}) = \left(\begin{bmatrix} -\frac{1}{8} & \frac{9}{8} \end{bmatrix}, \begin{bmatrix} \frac{3}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{3}{4} \end{bmatrix}, \begin{bmatrix} \frac{3}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{3}{4} \end{bmatrix} \right)$$
(13)

Moreover, the analysis of Section 4.3 reveals that the fiber $\phi_3^{-1}(\hat{p})$ consists only of the point $\hat{\theta}$ and the "swapped" point

$$\widehat{\theta}' = (\widehat{\pi}', \widehat{T}', \widehat{E}') = \left(\begin{bmatrix} \frac{9}{8} & -\frac{1}{8} \end{bmatrix}, \begin{bmatrix} \frac{3}{4} & \frac{1}{4} \\ \frac{1}{4} & \frac{3}{4} \end{bmatrix}, \begin{bmatrix} \frac{1}{4} & \frac{3}{4} \\ \frac{3}{4} & \frac{1}{4} \end{bmatrix} \right)$$
(14)

which is also not in Θ_{st} . Hence the image point $\widehat{p} = \phi_3(\widehat{\theta}) = \phi_3(\widehat{\theta}')$ is a non-negative point of $\overline{BHMM}(3)$ that does not lie in BHMM(3).

6.2. A semialgebraic model membership test

In light of the fact that not every nonnegative distribution in $\overline{\rm BHMM}(n)$ is in BHMM(n), the defining equations of $\overline{\rm BHMM}(n)$ are not sufficient to test a probability distribution for membership to the model. Using the method of Section 2.3, membership to HMM(2,k,n) can be tested by reducing to the k=2 to recover the parameters.

So, suppose we are given a distribution $p \in \Delta_p^{2^{n}-1}$ and asked to determine whether $p \in BHMM(n)$. The following procedure yields either

- (1) a proof by contradiction that $p \notin BHMM(n)$,
- (2) a parameter vector $\theta \in \Theta_{st}$ such that $\phi_n(\theta) = p \in BHMM(n)$, or
- (3) a reduction of the question to whether p lies in one of the lower-dimensional submodels of BHMM(n) discussed in Section 4.3.

How to proceed from (3) is essentially the same as what follows, using the birational parametrizations of the respective submodels given in Section 4.3.

To begin, we let $p' = \mu_3^n(p) \in \Delta_p^{2^3-1}$, i.e. we marginalize p to the distribution p' it induces on the first three visible nodes. Note that if $p \in BHMM(n)$ then $p' \in BHMM(3)$. Observing the moments m_I of p', if any denominators in the formulae of Corollary 4 vanish, then we end in case (3).

Otherwise, we let $(a, b, c, u, v) = \psi_3^{-1}(p')$, choose v_0 to be either square root of v, and let $a_0 = a/v_0$, $c_0 = c/v_0$. If p were due to some BHM process, then by Theorem 3, these would be its parameters, up to a simultaneous sign change of (a_0, b_0, v_0) . With this in mind, we define $\theta = (\pi, T, E)$ using (9). If (π, T, E) are not non-negative stochastic matrices, then $p \notin \text{BHMM}(n)$ and we end in case (1). If they are, we compute $p'' = \phi_n(\theta)$, and if p = p'' then we end in case (2). Otherwise p must not have been in BHMM(n), so we end in case (1).

Note that since all the criteria in this test are algebraic equalities and inequalities, this procedure implicitly describes a semialgebraic characterization of BHMM(n) for all $n \geq 3$.

6.3. Identifiability of parameters

By a rational map on a possibly non-algebraic subset $\Theta \subseteq \mathbb{C}^k$, we mean any rational map on the Zariski closure of Θ , which will necessarily be defined as a function on a Zariski dense open subset of Θ . We define polynomial maps on Θ similarly.

Let $\phi: \Theta \to \mathbb{C}^n$ be an algebraic statistical model, where as usual we assume $\Theta \subseteq \mathbb{C}^k$ is Zariski dense, and therefore Zariski irreducible. A (rational) parameter of the model is any rational map $s: \Theta \to \mathbb{C}$. Such parameters form a field, $K \simeq \operatorname{Frac}(\mathbb{C}^k)$. In applications such as [14], it is important to know to what extent a parameter can be identified from observational data alone. In other words, given $\phi(\theta)$, what can we say about $s(\theta)$? This leads to several different notions of parameter identifiability, as discussed by Sullivant, Garcia-Puente, and Spielvogel [22].

Definition 4. We say that a rational parameter $s \in K$ is

- (set-theoretically) identifiable if $s = \sigma \circ \phi$ for some set-theoretic function $\sigma : \phi(\Theta) \to \mathbb{C}$. In other words, for all $\theta, \theta' \in \Theta$, if $\phi(\theta) = \phi(\theta')$ then $s(\theta) = s(\theta')$.
- rationally identifiable if $s = \sigma \circ \phi$ for some rational map $\sigma : \phi(\Theta) \to \mathbb{C}$ (this notion is used without a name by Sullivant et al. [22]).

- generically identifiable if there is a (relatively) Zariski dense open subset $U \subseteq \Theta$ such that $s|_{U} = \sigma \circ \phi|_{U}$ for some set-theoretic function $\sigma : \phi(U) \to \mathbb{C}$.
- algebraically identifiable if there is a polynomial function $g(p,q) := \sum_i g_i(p_1, \dots, p_n)q^i$ on $\phi(\Theta) \times \mathbb{C}$ of degree d > 0 in q (so that g_d is not identically 0 on $\phi(\Theta)$) such that $g(\phi(\theta), s(\theta)) = 0$ for all $\theta \in \Theta$ (and hence all $\theta \in \mathbb{C}^k$).

Question 6.1. What combinations of BHM parameters are rationally identifiable, generically identifiable, or algebraically identifiable?

To answer this question we introduce a lemma on algebraic statistical models in general:

Lemma 2. For any algebraic statistical model ϕ as above, the sets K_{ri} , K_{gi} , and K_{ai} , of rationally, generically, and algebraically identifiable parameters, respectively, are all fields.

Proof. Since Θ is Zariski irreducible, so is $\phi(\Theta)$. Hence the set of rational maps on $\phi(\Theta)$ is simply the fraction field of its Zariski closure (an irreducible variety), and K_{ri} is the image of this field under $\phi^{\#}$, which must be a field.

For K_{gi} , the crux is to show that if $s, s' \in K_{gi}$ and $s \neq 0$ then $s'/s \in K_{gi}$. Let $U \subseteq \Theta$ and $\sigma : \phi(U) \to \mathbb{C}$ be as in the definition for s, and likewise $U' \subseteq \Theta$ and $\sigma : \phi(U') \to \mathbb{C}$ for s'. Let $U'' = \{\theta \in U \cap U' \mid s(\theta) \neq 0\}$, which, being an intersection of three Zariski dense open subsets of Θ , is a dense open. We have $\sigma \neq 0$ on $\phi(U'') \subseteq \phi(U) \cap \phi(U')$, so we can let $\sigma'' = \sigma'/\sigma : \phi(U'') \to \mathbb{C}$, and then $\sigma'' \circ \phi = s'/s$, so $s'/s \in K_{gi}$. Thus K_{gi} is stable under division, and simpler arguments show it is stable stable under +, -, and \cdot , so it is a field.

Finally, K_{ai} is expressly the relative algebraic closure in K of the image under $\phi^{\#}$ of the coordinate ring of $\phi(\Theta)$, which is therefore a field.

Proposition 14. For any algebraic statistical model ϕ as above, $K_{ri} \subseteq K_{gi} \subseteq K_{ai} \subseteq K$.

Proof. This is now just a restatement of Proposition 3 in [22].

Now, the answer to our identifiability question for BHM parameters can be given easily in the coordinates of Section 4. Here ϕ is the BHM map ϕ_n . The field K_{ri} is simply the image $\mathfrak{q}^{\#}(\operatorname{Frac}(\Theta'_{\mathbb{C}}))$ because by Theorem 2,

$$\psi^{\#}:\operatorname{Frac}(\overline{\operatorname{BHMM}}(n))\to\operatorname{Frac}(\Theta'_{\mathbb{C}})$$

is an isomorphism. Hence the rationally identifiable parameters are precisely the field of rational functions in $(a, b, c, u, v) = (a_0v_0, b, c_0v_0, u, v_0^2)$ (see (9) for the meanings of these parameters). Since K is a quadratic field extension of K_{ri} given by adjoining $v_0 = \sqrt{v}$, and K_{ai} is the algebraic closure of K_{ri} in K (almost by definition), it follows that $K_{ai} = K$, i.e. all parameters are algebraically identifiable. Finally, we observe that, by the action of sw in Section 4.1, there are generically two possible values of $v_0 = \frac{1}{2}(E_{11} - E_{01})$ for a given observed distribution, namely $\pm \sqrt{v}$. Hence $v_0 \notin K_{gi}$, and since a quadratic field extension has no intermediate extensions, it follows that $K_{ri} = K_{gi}$, i.e. all generically identifiable parameters are in fact rationally identifiable. In summary,

Proposition 15. For BHMM(n) where $n \geq 3$,

$$\mathbb{C}(a, b, c, u, v) = K_{ri} = K_{qi} \subseteq K_{ai} = \mathbb{C}(a_0, b, c_0, u, v_0)$$

6.4. A new grading on BHMM invariants

The re-parametrized model map ψ_n is homogeneous in cumulant and moment coordinates, with respect to a \mathbb{Z} -grading where $deg(m_v) = deg(k_v) = sum(v)$, deg(b) = 0, deg(a) = deg(c) = deg(u) = 1, and deg(v) = 2. This grading allows for fast linear algebra techniques that solve for low degree model invariants as in [5], except that this grading is intrinsic to the model. Bray and Morton's grading, which is in *probability* coordinates, is not on the binary HMM proper, but on a larger variety obtained by relaxing the parameter constraints that the transition and emission matrix row sums are 1. The invariants obtained in their search are hence invariants of this larger variety, and exclude some invariants of BHMM(n). The grading presented here can thus be used to complete their search for invariants up to any finite degree.

6.5. Equilibrium BHM processes

In Section 4.3 we found that if a BHM process is at equilibrium, our formula for ψ_3^{-1} is undefined. We may define Equilibrium Binary Hidden Markov Models, EBHMMs, by restricting ϕ_n to the locus $\{a_0b - a_0 + c_0 = 0\}$, which turns out to yield a four-dimensional submodel of BHMM(n) for each $n \geq 3$. The same techniques used here to study BHMMs have revealed that the EBHMMs, too, have birational parametrizations, and the ideal of EBHMM(3) is generated by the equations $m_1 = m_2 = m_3$ and $m_{12} = m_{13}$. The geometry of EBHMMs will need to be considered explicitly in future work to identify the learning coefficients of BHMM fibers.

6.6. Larger hidden Markov models

As we have remarked throughout, many results on BHMM(n) can be readily applied to HMM(2, k, n), i.e. HMMs with two hidden states and k visible states $\alpha_1, \ldots, \alpha_k$. For example, consider the parameter identification problem. We may specify the process by a $2 \times k$ matrix E of emission probabilities, along with a triple (a_0, b, c_0) defining the π and T of the two-state hidden Markov chain as in (9). As in Section 2.3, to obtain $E_{0\ell}$ and $E_{1\ell}$ from the observed probability distribution for any fixed j, we simply define a BHM process by letting $\alpha_{\ell} = 1$ and $\alpha_{j} = 0$ for $j \neq \ell$. Applying Proposition 6 to the moments of the distribution yields values for (a, b, c, u, v) provided the genericity condition that the denominators involved do not vanish. Letting $v_0 = \sqrt{v}$, $a_0 = a/v_0$, and $c_0 = c/v_0$, we obtain (a_0, b, c_0, u, v_0) up to a simultaneous sign change on (a_0, c_0, v_0) corresponding to swapping the hidden alphabet as in Section 4.1. Then $E_{0\ell} = u - v$ and $E_{1\ell} = u + v$, and we get π , T as well from (a_0, b, c_0) . We can repeat this for each $\ell = 1, \ldots, k$ to obtain all the emission parameters, and hence identify all the process parameters modulo the swapping operation.

REFERENCES 28

For each $\ell \in \{1, ..., k\}$, we can also obtain many polynomial invariants of HMM(2, n, k) by reducing to BHMM(n) as above, and marginalizing to collections of 4 equally spaced visible nodes to obtain points of BHMM(4) at which we know the invariants of Theorem 1 will vanish.

Given these extensions, one can hope that techniques similar to those used here could elucidate the algebraic statistics and geometry of HMMs with any number of hidden states as well.

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References

- [1] J. Baker. The DRAGON system An overview. Acoustics, Speech and Signal Processing, IEEE Transactions on, 23(1):24 29, Feb 1975.
- [2] Daniel J. Bates, Jonathan D. Hauenstein, Andrew J. Sommese, and Charles W. Wampler. Bertini: Software for numerical algebraic geometry. Available at http://www.nd.edu/~sommese/bertini.
- [3] Daniel J. Bates, Jonathan D. Hauenstein, Chris Peterson, and Andrew J. Sommese. Numerical decomposition of the rank-deficiency set of a matrix of multivariate polynomials, pages 55–77. Texts and Monographs in Symbolic Computation. Spinger-Verlag, 2010.
- [4] L.E. Baum and T. Petrie. Statistical inference for probabilistic functions of finite state Markov chains. *Ann. Math. Stat.*, 37:1554–1563, 1966.
- [5] Nicholas Bray and Jason Morton. Equations defining hidden Markov models. In *Algebraic Statistics for Computational Biology*, chapter 11. Cambridge University Press, 2005.
- [6] M. Casanellas and J. Fernandez-Sanchez. Performance of a new invariants method on homogeneous and non-homogeneous quartet trees. *Molecular Biology and Evolution*, 24:288–293, 2006.
- [7] David A. Cox, John B. Little, and Don O'Shea. *Ideals, Varieties, and Algorithms, Third Edition*. Springer New York, 2007.
- [8] Vesselin Drensky. Computing with matrix invariants. *Math. Balkanica*, 21:141–172, 2007.

REFERENCES 29

[9] Mathias Drton, Bernd Sturmfels, and Seth Sullivant. *Lectures on algebraic statistics*. Oberwolfach Seminars 39, 2009.

- [10] Nicholas Eriksson. Using invariants for phylogenetic tree construction. In *Emerging Applications of Algebraic Geometry*. I.M.A. Volumes in Mathematics and its Applications, 2008.
- [11] Daniel R. Grayson and Michael E. Stillman. Macaulay2, a software system for research in algebraic geometry. Available at http://www.math.uiuc.edu/Macaulay2/.
- [12] Alexandre Grothendieck and Jean Dieudonné. Éléments de géométrie algébrique (rédigés avec la collaboration de Jean Dieudonné) IV. Étude locale des schémas et des morphismes de schémas, Troisième partie. Publications Mathmatiques de l'IHÉS, 1966.
- [13] Anders Krogh, I. Saira Mian, and David Haussler. A Hidden Markov Model that finds genes in E. coli DNA. *Nucleic Acids Research*, pages 4768–4778, 1994.
- [14] Nicolette Meshkat, Marisa Eisenberg, and Joseph J. DiStefano. An algorithm for finding globally identifiable parameter combinations of nonlinear ode models using Gröbner bases. *Mathematical Biosciences*, 222(2):61 72, 2009.
- [15] Lior Pachter and Bernd Sturmfels. Algebraic Statistics for Computational Biology. Cambridge University Press, 2005.
- [16] Giovanni Pistone, Eva Riccomagno, and Henry P. Wynn. Computational commutative algebra in discrete statistics. Chapman and Hall / CRC, 2001.
- [17] Alexander Schönhuth. Generic identification of binary-valued hidden Markov processes. arXiv:1101.3712, 2011.
- [18] K.S. Sibirskii. Algebraic invariants for a set of matrices. Siberian Mathematical Journal, 9:115–124, 1968. ISSN 0037-4466.
- [19] Jim Q. Smith and Piotr Zwiernik. Tree cumulants and the geometry of binary tree models, 2010. arXiv:1004.4360v3.
- [20] Ruslan L. Stratonovich. Conditional Markov Processes. Theory of Probability and its Applications, pages 156–178, 1960.
- [21] Bernd Sturmfels and Piotr Zwiernik. Binary cumulant varieties, 2011. arXiv:1103.0153.
- [22] Seth Sullivant, Luis David Garcia-Puente, and Sarah Spielvogel. Identifying causal effects with computer algebra. Proceedings of the 26th Conference of Uncertainty in Artificial Intelligence, 2010.

REFERENCES 30

[23] Sumio Watanabe. Algebraic Geometry and Statistical Learning Theory (Cambridge Monographs on Applied and Computational Mathematics). Cambridge University Press, 2009.

[24] Byung-Jun Yoon. Hidden Markov Models and their Applications in Biological Sequence Analysis. $Current\ Genomics,\ 10(6):402-415,\ 2009.$